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ABSTRACT

Prepared cooperatively by the Bureau of the Census and the Economic Research Service of the U.S. Department of Agriculture, this document presents narrative and tabular data on: demographic and social characteristics of the farm population; economic characteristics of the farm population; revision of farm population processing procedures; and related reports. Tables within the text include: population of the United States, total and farm: April 1960 to 1976; fertility characteristics of farm and nonfarm women, by race: 1976; metropolitan-nonmetropolitan residence of the farm and nonfarm population, by race: 1976; persons 14 years old and over employed in agriculture, by farm-nonfarm residence and sex: April 1976 and 1970; nonfarm residents 14 years old and over employed in agriculture, by class of worker and sex: April 1976 and 1970; and income characteristics of farm and nonfarm families by race: 1975. Detailed tables are presented as follows: farm population, by race and sex, for broad age groups: April 1976 and 1970; farm population, by age and sex: April 1976 and 1970; characteristics of farm and nonfarm families, by race: 1976; employment status of the farm population 14 years old and over, by sex, April 1976 and 1970, and by region, April 1976; farm residents 14 years old and over employed in . agriculture, by class of worker, race, and sex, April 1976 and 1970, and by region: April 1976; farm residents 14 years old and over employed in nonagricultural industries, by class of worker, race, and sex for regions: April 1976. (JC)

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CURRENT POPULATION REPORTS

Farm Population

P-27, No. 49 Issued December 1977

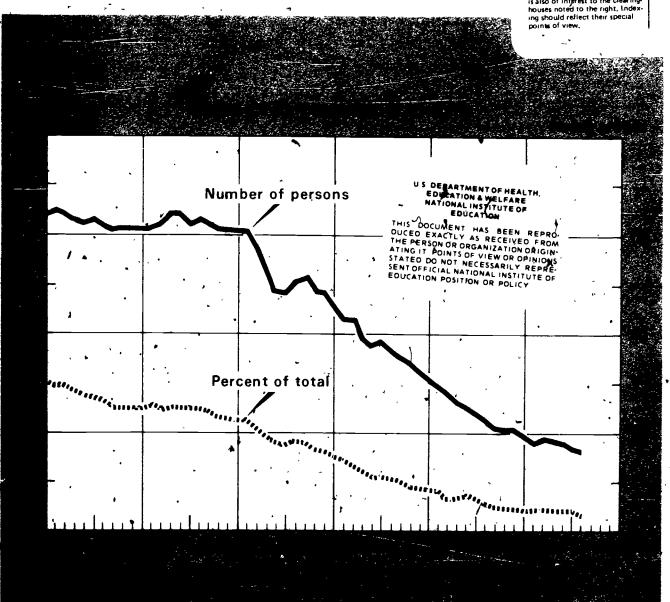
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Farm Population of the United States: 1976

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This report was prepared jointly by Vera J. Banks, Economic Research Service, U.S. Department of Agriculture, and Diana DeAre and Robert C. Speaker, Population Division, U.S. Bureau of the Census:

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CURRENT POPULATION REPORTS

Farm Population

Series Census—ERS P-27, No. 49 Issued December 1977

Farm Population of the United States: 1976

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SYMBOLS USED IN TABLES

- Represents zero or rounds to zero.

B Base less than 75,000.

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Farm Population of the United States: 1976

(Advance data on the 1976 farm population were usued in April 1977 in Current Population Beports Series P-27, No. 48)

The number of persons living on farms in rural areas of the United States averaged 8,253,000 for the 12-month period centered on April 1976. About 1 person out of every 26, or 3.9 percent of the Nation's 214 million people, had a farm residence (table A). These estimates were prepared cooperatively by the Bureau of the Census and the Economic Research Service, U.S. Department of Agriculture

The farm share of the total U.S. population has declined fairly steadily over the last 55 years (figure 1) in 1920, when the farm population was first enumerated separately, 30 percent of the Nation's population resided on farms. The proportion had fallen to 15 percent by 1950, to 5 percent by 1970, and has now dropped below 4 percent.

Although the overall trend has been one of decline, the rate of decrease in the farm population has shown short-term fluctuations. After declining at an average annual rate of 4.8 percent during the 1960's, the farm population appeared to be leveling off in the early 1970's. The average rate of decline from 1970 to 1974 was 1.2 percent per year. Since 1974, the rate of loss has accelerated, with an average annual decline from 1974 to 1976 of 5.8 percent. It should be noted, however, that this rate is samewhat inflated because of changes made in 1976 involving the procedures for processing survey information on farm and nonfarm residence. (These changes are explained in detail in a later section of this report.) Approximately 130,000 of the 611,000 decline in the farm population between 1975 and 1976 may be attributable to processing changes.

DEMOGRAPHIC AND SOCIAL CHARACTERISTICS OF THE FARM POPULATION

Distribution. While the farm population is primarily non-metroplitan, one-fifth of the farm total lives within the boundaries of standard metropolitan statistical areas. (SMSA's) as defined in 1970 (table B). In comparison, nearly 70 percent of the nonfarm population lives within the 1970 metropolitan boundaries. Data for families from the March.

Table A. Population of the United States, Total and Farm: April 1960 to 1976

(Numbers in thousands)

	- 1		н -
	Total	Farm po	opulation
Year	resident population	Number	Percent of total
	· , ·	persons ¹	population
- ·	•		1
1976	214,284	8,25 3	3.9
1975	212,542	8,864	4.2
1974	211,018	9,264	4.4
1973	209,468	9,472	4.5
1972	207,802	9,610	4.6
1971	205,677	9,425	4.6
1970	² 203, 235	9,712	8
1969	200,887	10,307	5.1
1968	198,923	10,454	5.3
1967	196, 976	10,875	5.5
1966	195,045	11,595	5.9
1965	192,983	• 12,363	6.4
1964	190,507	12,954	1. 6.8
1963	187,837	13,367	7.1
1962	. 185,104	14, 313	1 .7.7
1961	182,298	14,803	8.1
1960	²1,79, 3 23	15, 635	8.7

¹Five-quarter averages centered on April; see "Definitions and Explanations." ²Official census count.

1976 Current Population Survey (CPS) indicate that the majority of the metropolitan farm population resides in SMSA's of less than 1 million population.²

There is a difference by race in the distribution of the farm population by metropolitan nonmetropolitan residence.



[&]quot;Estimates of the farm population from 1920 to the present are not strictly comparable due to definitional changes. Prior to 1960, farm residence was based essentially on self-identification, i.e., respondents themselves determined whether or not they used on a farm. From 1960 through 1976, the farm population has been restricted to persons living in rural territory and has been identified on the basis of acreage and sales information (see "Definitions and Explanations" in the appendix).

² U.S. Bureau of the Census, Current Population Reports, Series P-20, No. 311, "Household and Family Characteristics March 1976." Much of this farm population is "metropolitan" in little more than a technical sense, being included in SMSA's because the latter are defined in terms of entire counties and thus frequently include non-suburbanized territory. The metropolitan farm population may have a certain significance, however, as representing farm residents who live close to sizable cities.

2

Black³ farm residents are more likely to live in nonmetropolitar areas than are White. Approximately 90 percent of the Black farm population resided in nonmetropolitan areas. In 1976; the comparable proportion for Whites was 80 percent. In contrast, Blacks who live off farms are more likely to, be in metropolitan areas than are their White counterparts

Racial composition. The farm population is becoming increasingly. White due to differences in the rates of decline of Whites and Blacks (table 1). Blacks on farms numbered

541,000 in 1976 and represented percent of the total farm population; the corresponding proportions in 1960 and 1970 were 16 and 10 percent, respectively. Between 1970 and 1976, the number of Whites on farms decreased by 12 percent as compared with a 42 percent decrease for Blacks. The annual rates of loss over this 6 year period averaged 2.2 percent for Whites and 9.2 percent for Blacks. A similar differential in rates of decline was experienced over the preceding decade. From 1960 to 1970, the number of Whites on farms decreased at an average annual rate of 4.0 percent, while the a grage rate of loss for Blacks was 10.1 percent per year.

Historically, higher rates of population loss among Black farm residents have been associated with heavy losses in the number of cotton and tobacco tenant farmers. Blacks have had a disproportionate representation among tenant farmers,

Table B. Metropolitan Normetropolitan Residence of the Farm and Nonfarm Population, by Race: 1976

(Figures are five-quarter averages centered on April)

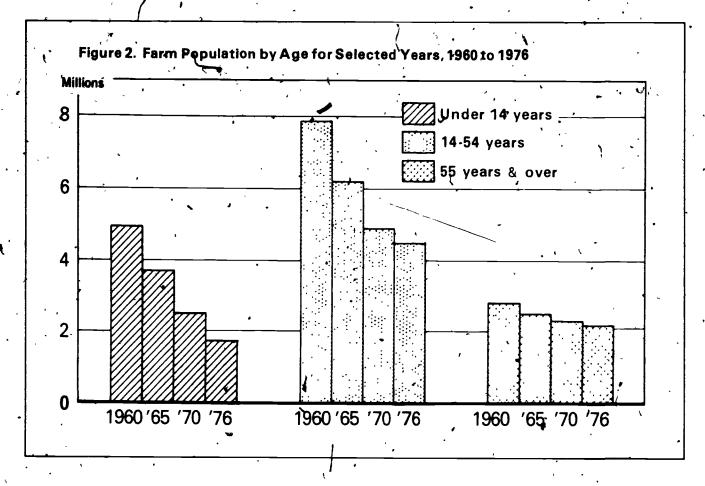
Page and modification	Total	Farm	Nonfarm
Race and residence			
ALL RACES	•		
United Statesthousands Inside SMSA's ² thousands Percent Outside SMSA'sthousands Percent	1210,332	8,253	202,079
	142,567	1,560	141,007
	67.8	18.9	69.8
	67,765	6,698	61,072
	32.2	81.1	30.2
United States	182,638	7,711	174,927
	121,392	1,498	119,894
	66.5	19.4	68.5
	61,246	6,213	55,033
	33.5	80.6	31.5
United Statesthousands Inside SMSA'sthousands Percentthousands Outside SMSA'sthousands	27,694	541	27,153
	21,175	62	21,113
	76.5	11.5	77.8
	6,519	479	6,040
	23.5	88.5	22.2

The total U.S. population figure shown here differs from that shown in table A because table A refers to the total resident population, whereas this and other tables refer only to the civilian populational population.

²SMSA's refers to standard metropolitan statistical areas as designate in the 1970 census publications; see "Definitions and Explanations."

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The data for Blacks in the text refer to Blacks and persons of races other, than White. In the 1970 census, Blacks comprised 90 percent of the total population other than White and 87 percent of the farm population other than White.



and the number of such farms has fallen steadily and sharply since 1935. With mechanization and modernization of cotton and tobacco farming, landowners have, for the most part, ceased to employ tenant labor to produce their crops. Declines in the number of small farms and of hired workers who live on farms have also contributed to the disproportionate drop in the Black farm population.⁴

Age. The farm population has been characterized for many years by unequal rates of population loss between the two broad age groups—under 14 years and 14 years old and over. Since 1970 the number of farm children under 14 years of age has dropped by a third, and their proportion of all farm people has declined from 26 to 20 percent (figure 2 and table 2). During this same period, the number of farm persons 14 years old and over decreased by only 9 percent. This decline in children reflects both the high net outmigration in earlier decades of young farm adults of childbearing age and the sharp drop in the national birth rate in the early 1970's, which extended to farm as well as nonfarm areas.

Higher rates of decrease among those under 14 years of age characterized both the White and Black farm populations. During the 1970 to 1976 period, the number of White children on farms declined by 28 percent, while the White adult farm population decreased by only 7 percent (table 1). The comparable rates of decline for Blacks were 60 and 32 percent, respectively. Despite the marked difference in rates of population loss, there is some evidence that children continue to comprise a greater proportion of the Black farm population than they do of the White farm population. In 1976, 25 percent of all Blacks on farms were under 14 years of age, compared with 20 percent of White farm residents.

The pattern of decline was not consistent for different age groups of adults. Over the 1970 to 1976 period, young adults aged 20 to 34 years rose as a proportion of the total farm population from 13 to 16 percent. No significant changes occurred in the proportions of farm teenagers—those 14 to 19 years old—or of the older age categories of farm adults.

Sex. The dwindling size and changing age structure of the farm population has not affected the continuance of another of its distinctive features—more males than females. Farm males outnumbered farm females by 358,000 in 1976; there were 109 males on farms for every 100 females (table 2). In comparison, there were only 93 males per 100 females in the nonfarm civilian noninstitutional population. The stronger representation of males in the farm population reflects a

⁴ Vera J. Banks and Column*L. Beale, "Farm Population by Race, Tenure, and Economic Scale of Farming, 1960 and 1970," Agricultural Economic Report No. 228, U.S. Department of Agriculture, Economic Research Service, 1972; and Calvin L. Beale, "The Black American in Agriculture" in Mabel M. Smythe, ed., The Black American Reference Book (Englewood Cliffs, N.J. Prentice-Hall, 1976).

somewhat higher rate of outmigration of females as compared with males. This outmigration of females from farms, typically as they reach maturity, reflects the predominately masculine nature of farm work; of the 2 million farm residents amployed in agriculture in 1976, 16 million or more than four-fifths were male (table D)

Family type and size. Data from the March 1976 CPS indicate that a greater proportion of farm families than of nonfarm families have both husband and wife present (table 3). While 92 percent of farm families include both husband and wife, the comparable figure for nonfarm families is 84 percent. This difference between farm and nonfarm families exists for both Whites and Blacks.

The average sizes of farm and nonfarm families in March 1976 were 3.5 and 3.4 persons, respectively—a tifference which is not statistically significant. Nor was there any significant difference between the average size of White farm families (3.4 persons) and White nonfarm families (3.3 persons). Among Blacks, however, there is some statistical evidence that the average of 4.6 persons per farm family was higher than the 3.9 person average for nonfarm families.

While the average sizes of farm and nonfarm families are not significantly different, the distributions of families by number of persons show that large families—those with 6 or

Dale E Hathaway, J Allen Beegle, and W Feith Bryant, People of Rural America, U.S. Bureau of the Census, 1960 Census Monograph (Washington, D.C. U.S. Government Printing Office, 1968), pp. 68-71, and Henry S. Shryock, Jacob S. Siegel, and Associates, The Methods and Materials of Demography, U.S. Bureau of the Census (Washington, D.C. U.S. Government Printing Office, 1971), pp. 193-199

percent) than of nonfarm families (9 percent). The higher proportion of large families within the farm population is partially due to the presence, among families which have children, of a greater number of children within farm families. Among families with own children under 18 present, 9 percent of farm families have 5 or more children compared to only 4 percent of nonfarm families. This difference is not completely reflected in the mean family size estimates because of the offsetting effect of the smaller proportion of farm families with own children under 18 present, 47 percent of farm families have own children under 18 compared to 54 percent of nonfarm families

Fertility. The fertility of farm women continues to be higher than that of nonfarm women. Data for June 1976 (table C) indicate that the average number of children born to farm women 15 to 44 years of age who have ever been married (2,699 per 1,000 women) is significantly higher than the average born to nonfarm women of comparable age (2,064 per 1,000 women). However, there is some evidence that the difference is mainly attributable to the fertility experience of women in the oldest (35 to 44 years) age group, who have essentially completed their childbearing.

Table C also presents June 1976 data on birth expectations of currently married women 14 to 39 years of age. Farm women in this age group expected to have a lifetime total of 2,947 births per 1,000 women. Although this figure is significantly higher than the 2,428 births per 1,000 expected by nonfarm women, it should be noted that this difference in lifetime births expected is due entirely to a difference in the number of births to date.

Table C. Fertility Characteristics of Farm and Nonfarm Women, by Race: 1976

	For m	eaning	d £ sy mb	ols, se	e text)	\		, 		
	,A	11 race	s	White			Black a	nd other	r ŕaces	
Characteristic	Potal	Farm	Non- farm	Total	Farm	Non- farm	Total	Farm	Non- farm	
Children ever born per 1,000 women ever married: Total, 15 to 44 years 15 to 24 years	2,082 837 1,892 3,058	2,699 909 2,123 3,514	. *836	2;017 777 1,848 2,979	2,699 812 2,100 3,528	1,955 777 1,842 2,952	2,552 1,319 2,204 3,588	(B) (B) (B)	2,550 1,310 2,200 3,591	
Married women 14 to 39 years old¹: Births to datc per 1,000 women	1,876	2,500 2,947	1,859	1,842 2,415	2,472 2,914	ľ	2,211 2,708	(B)	2,199 2,696	

Data limited to currently married women reporting on birth expectations.

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Source: Unpublished data from the June 1976 Current Population Survey. See table A-6 for bases and table A-4 for standard errors.

ECONOMIC CHARACTERISTICS OF THE FARM POPULATION

Labor force participation. In 1976, about 4 million persons, or three fifths of the farm population 14 years old and over, were in the labor force, either employed or seeking work (table 4). Although there has been some decline in total number, the rate of labor force participation among farm residents has remained essentially unchanged since 1970

Overafi, the level of labor force participation among farm residents was about the same as that among nonfarm residents. However, there were significant differences between these two residence groups by sex. Farm males were more likely to be in the labor force than nonfarm males, in 1976, the labor, force participation rates for the two groups were about 80 percent and 74 percent, respectively. On the other hand, the level of labor force participation of farm women was below that of their nonfarm counterparts. About 40 percent of all females 14 years old and over living on farms were either working or looking for a job in 1976. In comparison, females living off farms had a labor force participation rate of 46 percent.

Farm males also exhibited some differences in labor force participation by region of residence. Males 14 years old and over living on farms in the combined Northern and Western States were more likely than southern farm residents to be in the labor force (82 percent versus 76 percent). Among the female farm population 14 years old and over, about two-fifths were in the labor force irrespective of region.

As in earlier years, labor force participation was somewhat higher among White farm residents than among Black farm residents. In 1976, the labor force participation rates for these two racial groups were 61 and 54 percent, respectively (table 5). This racial disparity in labor force participation is accounted for by, differences in the participation of males, as there was no significant difference by race in the likelihood of females being in the labor force. In the male farm population 14 years old and over, the rate of labor force participation was 80 percent for Whites and 67 percent for Blacks, among females, both races had rates of 40 percent.

Agricultural and nonagricultural employment. In 1976, 2.0 million persons, or 51 percent of the employed farm resident labor force, were engaged solely or primarily in agriculture (table 4). This represents a decline of 16 percent, or about 370,000 workers, in primary agricultural employment among farm people since 1970. During this same 1970 to 1976 period, there was no significant change in the number of farm residents working in nonagricultural industries about 1.9 million workers. However, nonagricultural employment as a proportion of total employment of farm people rose from 45 to 49 percent. Figure 3 shows that as a consequence of the trends in these two proportions, the farm resident labor force in 1976 was almost equally divided between employment in agriculture and employment in nonagricultural pursuits.

There is some evidence that the proportion of the farm resident labor force employed in nonagricultural industries has risen in both of the major regions of the country since 1970. However, southern farm residents are more likely to be employed in nonfarm work than are farm residents of the combined North and West. In 1976, 57 percent of employed southern farm residents worked in nonfarm jobs; outside the South, the proportion was 45 percent. This disparity is apparently associated with the relatively large number of low-income farms in the South, whose residents sought supplemental nonfarm income. Preliminary data from the latest Census of Agriculture (1974) reveal that Southern States contain two-fifths of all farms in the United States but nearly three-fifths of those with sales of less than \$2,500.6

Employment in nonagricultural industries was more prevalent among farm females than among farm males. In 1976, about 7 out of 10 employed farm resident women were engaged in nonagricultural pursuits, among farm resident males, only 4 out of 10 were so employed

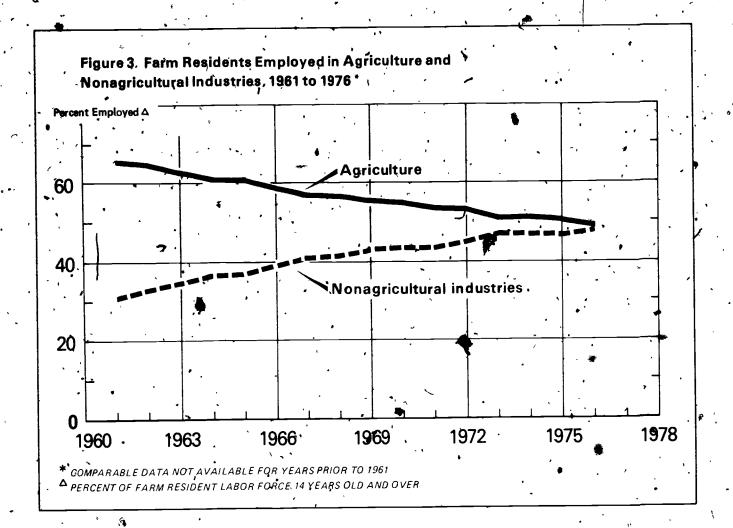
Unemployment. The rate of unemployment—the proportion of the civilian labor force currently without a job and looking for work—was relatively low in the farm population. In 1976, 2.7 percent of the labor force living on farms was unemployed (table 4). The comparable rate for the civilian noninstitutional population living off farms was 8.2 percent Within the farm resident labor force, unemployment was higher among Blacks than among Whites, the rates of unemployment in 1976 for these two racial groups were 7.3 percent and 2.4 percent, respectively (table 5). However, despite this difference, the rate of farm unemployment was below that of the nonfarm population for each racial group. For the civilian noninstitutional population living off farms, the rates of unemployment averaged 13.8 percent for Blacks and 7.5 percent for Whites over the period covered by the survey.

The frequency of holding two or more jobs among persons employed in agriculture is thought to contribute to lower unemployment among farm residents. In May 1976, over 800,000 multiple jobholders, one fifth of the total, had at least one job in agriculture. Most of this latter group were nonagricultural wage and salary workers who operated their own farms as a secondary job. Thus, farm operators with qual employment who lose their nonfarm job are not considered as unemployed because of their continued employment in farm work.

Class of worker. Of the 20 million farm residents employed in agriculture in 1976, 62 percent were self-employed and 19 percent were employed in each of the remaining two classes—wage and salary workers and unpaid family workers (table 6) Self-employment was the major class of work among farm—persons employed in agriculture irrespective of



⁶U.S. Department of Commerce, Bureau of the Census 1974 Census of Agriculture. Preliminary Reports Washington, D.C. 1976 ⁷U.S. Department of Labor, Bureau of Labor Statistics, Multiple Jobholders, May 1976, Special Labor Force Report 194 1977.



region of residence. However, there were significant differences in the class of worker distribution by sex. Self-employment was the dominant class of work among farm males (70 percent), while farm females were most often unpaid family workers (63 percent)

A difference in class of worker distribution was also apparent by race. In 1976, self-employment was the dominant class of work for 63 percent of White farm residents employed in agriculture, whereas only 30 percent of Black farm residents were self-employed. This lower incidence of self-employment (a category that consists chiefly of farm operators) reflects the comparatively small number of farms operated by Blacks. The Census of Agriculture indicates that less than 5 percent of all U.S. farms have a Brack operator.

The indicated decline between 1970 and 1976 in the number of farm residents employed in agriculture occurred only among self-employed workers and unpaid family workers. During this 6-year period, there was no significant decrease in the number of farm resident wage and salary agricultural workers.

in 1976, there were 1.9 million persons living on farms and working in nonagricultural industries. As in earlier years,

these farm resident nonagricultural workers were predominantly wage and salary workers regardless of race, sex, or, region of residence (table 7)

The total number of persons employed solely or primarily in agriculture in the United States averaged 3.6 million in 1976 (table D). Of these, a little more than half (55 percent) lived on farms, while the remainder lived off farms and commuted to work. Although the data for 1976 and 1970 imply a decrease of about 100,000 in total agricultural employment, the estimated decline is not statistically significant. Although there has been no significant change in the total number of agricultural workers, there has been an increase in both the number and proportion of agricultural workers with a nonfarm residence. Between 4970 and 1976 the number of , nonfarm resident agricultural workers rose from 1.4 million to 1.6 million; and their proportion of the total increased from 37 percent to 46 percent. This reflects the increasing trend among farm wageworkers to commute from nonfarm residences to their farm jobs. In 1976, about three out of every#four wage and salary agricultural_workers lived off farms (see tables E and 6). In contrast, self-employed and ungaid family workers in agriculture continue to be mainly farm residents.

Unlike their farm counterparts, who, as discussed earlier, show variations in the class of worker distribution by sex and race, nonfarm resident agricultural workers are primarily wage and salary workers regardless of sex or race.

income. The median income of farm families was \$10,845 in 1975, substantially lower than the \$13,829 for inflarm families (table F). Although this represents a difference of nearly \$3,000, the gap is only two thirds the farm nonfarm income differential that existed in 1970. Farm median family, income in 1970 was \$4,500 less, in terms of 1975 dollars, than that of nonfarm families. Since 1970, the median income of farm families has increased by 15 percent, while that of nonfarm families has shown no significant change in real terms.

The contrast between farm and nonfarm income levels is particularly sharp among Black families. While Black nonfarm median family income was \$9,404 in 1975, Black farm median family income was only \$4,857. The latter figure also presents a striking contrast to that of White farm families of \$11,237), being only about two fifths as large.

The proportion of farm families who are below the wincome level (13.7 percent) is higher than that of nonfarm families (9.5 percent). Among Blacks, the proportion of farm

families below the low-income level is 54 percent, about 5½ times as high as the national average for all families and about 4½ times as high as that for White farm families.

REVISION OF FARM POPULATION PROCESSING PROCEDURES

In February 1976, two changes were made in the Cyrrent Population Sulvey (CPS) procedures for determining farminonfarm residence of the rural population. The first was necessitated by the change in the official farm definition announced in late 1975 (see "Definitions and Explanations" in the appendix). The questions asked of respondents were altered to enable collection of data under the new definition as well as continuation of collection under the previous definition. (A detailed description of the old and new questions is provided in the appendix.) Basically, the first change involved the addition of a greater number of farm sales intervals. The second change entailed a refinement in the procedure for imputing farm nonfarm residence for households

Table D. Persons 14 Years Old and Over Employed in Agriculture, by Farm-Nonfarm Residence and Sex: April 1976 and 1970

(Numbers in thousands. Figures are five-quarter averages centered on April)

	Both sexes Male			10	Female			Percent distribution				
Residence	Both	sexes male			remaie .		Both sexes		Male		Female	
	1976	1970	1976	1970	19.76	1970	1976	1970	1976	1970	1976	1970
Total employed in agriculture	3,592	3,696	2,941	3,045 2	651	65.0	ጉ 100.0	100.0	100.0	100.0	100.0	100.0
Farm residents	1,960 1,633	2,333 1,363	1,618 1,323		342 310	431 220	54.6 45.5	63,1		62.5	52.5 47.6	66.3 4 3.8

Table E. Nonfarm Residents 14 Years Old and Over Employed in Agriculture, by Class of Worker and Sex: April 1976 and 1970

(Numbers in thousands. Figures are five-quarter averages centered on April)

		, ,	Dath asses		Male		,		Percent distribution.						
	Class of worker	Both sexes		male		Female /		Both sexes		Male		Female			
•		1976	197,0	1976	1910	1976	1970	1976	1970 .	1976	1970	· 1976 ·	1970		
. —,	Total agricultural workers	1,633	1,363	1,323	1, 143	310	220	100.0	100.0	100.0	100.0	100.0	100.0		
Wage	employed workers and aslary workers	481 1,092 59	1	436 863 24	396 ,719 27	45 228 35	28 153 39	29.5 66.9 3.6	31 -1 64.0 4.8	33.0 65.2 1.8	34.6 62.9 2.4	14.5 73.5 11.3	12.7 69.5 *17.7		



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^{*} Alli data presented in this report refer to the previous farm

with nonresponses to the residence questions. These changes were made in an effort to reduce the nonsampling error in the estimates (i.e., certain response, enumeration, and processing errors). The differences which result from the addition of a greater number of sales intervals and the changed imputation procedure are due entirely to changes in the magnitude of nonsampling error associated with the statistics. The revised procedure is described below.

Change in imputation procedure. In the late 1960's, the Cersus Bureau implemented a simple imputation system to provide missing responses for those households failing to answer the acreage and farm sales questions used to determine farm-nonfarm residence in the CPS. The imputation approach assigns to a sample household with missing responses the information from a "similar" sample household that did respond to the questions. Thus, during the processing of the household records, any rural household with a missing response to either or both of the residence questions was assigned the residence classification (farm or nonfarm) of the last "good" rural household, i.e., the last processed record with answers reported for both residence items.

The revised CPS imputation system for missing acreage the farm sales information follows the same general procedure. A difference arises only when an answer is reported for one of the residence items (acreage or sales) while the other is left blank. In such a situation, the new procedure makes use of the reported item by imputing the missing response from the last "good" rural record with a similar response on the reported item. Earm or nonfarm residence under the old farm definition is then determined on the basis of the re-

ported residence item and the imputed item. Under the new farm definition, residence is determined on the basis of the reported, or imputed, farm sales item only:

The following is a simplified illustration of how the refined imputation system provides an estimate of acreage or farm sales for a household not responding to the specific questions on the CPS questionnaire:

A rural household reports that the place it owns or rents has 10 acres or more but does not provide the CPS interviewer with information on farm sales. The imputation system is designed to seek out the last rural mousehold that reported lacreage of 10 acres or more and a farm sales per sales experience to be sales were in the \$50-\$249 interval.

rew procedures are introduced into a data processing system, some changes in the results must be expected. Therefore, the data were examined to determine the effects, if any, the new, question design and imputation system had on the level, geographic distribution, and characteristics of the farm population. The analysis indicated that although the level of farm population was affected by the revisions, its demographic characteristics. In enot significantly altered.

Although it was not possible to separate the effects of the revised imputation procedure from the new question design, the condinated effect of these two procedural changes could be approximated. It was estimated that 130,000 of the

Table F. Income Characteristics of Farm and Nonfarm Families, by Race: 1975

		All races		; ;	White		Black and other races			
Characteristics	Total	Farm	Nonfarm	Total	Farm.	Nonfarm	Total	· Farm	Nonfarm	
Total familiestususands	56,245	. 2,200	54,045	49,873	2,105	47,768	6-373	95	6,277	
Families by 1975 income	100.0	100.0	100.0	100.0	100.0	100.0	100 0		100.0	
Less than \$4,000 or loss	8.0	. 14.8	7.7	. 6.6	13.5	. 6.3°		43.2	18.3	
\$4,000 to \$9,999	1	31.7	25.0	. 24.1	. 31.4	23.7	34.7	36.8	34.7	
\$10,000 to \$14,999		19.9	22.4	⁻ 22.6	′20.5	22:7	20.2	6.3	20.4	
\$15,000 to \$19,999	18.8	13.6	19.0	19.5	14.0	19.7	13.3	5.3	13.4	
\$20,000 and over	1	20.0	25.9	27.3	20.5	27.6	13.2	7.4	13.3	
***	<i>F</i>	<u> </u>	Ì	^.	ł		•	•	1	
Median family income	, ,		Ι.,	j					,	
(1975 dollars):	.,		1.	1,			40,000	*/ 057	\$9,404	
. 1975.	\$13,719	\$10,845		\$14,268	\$11,237		. \$9,321	\$4,857	9,433	
1974	14,081	11,582	14,179	14,634		14, 753		6,003		
1973		12,167	14, 719	15,254	12,570	15,391	9,200	5,536	9,301	
1972	14,300	11,389	14,447	14,858	11,722	15,025	. 9,142	5, 494	9,230	
1971		9,561	13,868	14, 182	9,851	14,398	8,922	4,921	9,053	
1970	13,676	9,393	13,876	14, 189	9,730	14,405	9,032	4,278	9,192	
Percent of families	100.0	100.0	100.0	100.0	100.0	100.0	100.0	100.0	1000	
Below low-income level		13.7	•	7.7	11.9	7.5	25.3	53.7	. 24.9	
Above low-income level	•1	86.3	90.5	92.3	-88-1	92.5	74.7	. 46.3	[75 .]	
WOOAS TOM-THEOMS TOAST !	1 ,0.3	1 00.3	1			L	L	l	J	

611,000 decline in farm population from 1975 to 1976 may be attributed to the new procedures. A comparative analysis of data before and after introduction of the new procedures revealed some variation but no significant difference in the regional distribution, age, race, sex, or employment characteristics of the farm population.

RELATED REPORTS

Comparable figures for 1976 appear in Farm Population, Series Census ERS (P-27), No. 47, and earlier reports were published annually beginning in 1961. Beginning with 1972, the data are not strictly comparable with data for earlier years because of adjustments in sample design and survey procedures occasioned by 1970 census data. However, the effect on comparability with prior data is not considered sufficient to warrant revisions of earlier statistics. Application of 1972 procedures to data for March 1970 lowered the farm population 14 years old and over by about 75,000.

Although not fully comparable with CPS, farm population figures for 1970 for the United States, States, and counties appear in chapter C of 1970 Census of Population, Volume I, Characteristics of the Population; characteristics of the farm population by States are presented in chapter D.

Table 1. FARM POPULATION, BY RACE AND SEX, FOR BROAD AGE GROUPS: APRIL 1976 AND 1970

(Numbers in thousands. Figures are five-quarter averages centered on April)

• •	Poth	Both sexes		Male, Fema		-1-		Pe	Percent distribution					
Age and race	Potii sexes				Female		Both sexes		Male		Female			
•	1976	1970	1976	1970	1976	1970	1976	1970	1976	1970	1976	1970		
Total	8,253	9-, 712	4,305	5,004	-3 ,947	4,708.	100.0	100.0	100.0	100.0	100.0	100.0		
White	7,711 541	8,7 7 5 938	4,026 279	. 4,524 480	3,685 262	4,251 458	93.4 6.6	90.4 9.7	93.5 6.5	90.4 9.6	93.3 6.6	90.3 9.7		
Under 14 years	1,676 1,541	2,490 2,152	878 812	1,274 1,101	797 - 729	1,216 1,051	100.0 91.9	100.0 86.4	100.0 92.5	100.0	100.0 91.5	100.0 86.4		
Black mid other races	134	338	66'	173	68	165	8.0	13.6	7.5	13.6	8.5	13.6		
14 years and over White	6,577 -6,170 407	7,222 6,623 600	3,427 3,214 213	3,730 3,423 307	3,150 2,956 194	3,492 3,200 · 293	100.0 93.8 6.2	100.0 91.7 8.3	100.0 93.8 6.2	100.0 91.8 8.2	100.0 93.8 6.2	100.0 91.6 8.4		

Table 2. FARM POPULATION, BY AGE AND SEX: APRIL 1976 AND 1970

(Numbers in thousands. Figures are five-quarter averages centered on April)

•	Dath		,				Female		Percent distribution							
Age .	Both .	sexes		ile .	rem	ale	Both	sexes	, Ma	le '	Fem	ale				
	1976	1970	1976	1970 .	1976	1970	1976	1970	1976	1970	1976	1970				
All ages	8,253	9,712	4,305	`5,004	3,947	4,708	100.0	1,00.0	100.0	100.0	100.0	100.0				
Under 14 years	1,676 6,577	2,490 7,222	. 878 3,427	1,274 3,730	797 3.150	1,216	20.3 79.7	25:6 74.4	20.4 79.6	25.5 74.5	20.2 79.8	25.8 74.2				
14 to 19 years	1,193 531	1,316 502	651 299	714 269	542 233	602 232	14.5 6.4	13.6 5.2	15.1 6.9	14.3 5.4	13.7 5.9					
25 to 34 years	755 912	770 1,061	" 396 439	371 -518	* 359 473	399 543	9.1 11.1	7.9 10.9	9.2 10.2	7.4	9.1 -12.0	8.5 11.5				
45 to 54 years	1,127 1,070 989	1,250 1,202 1,122	579 559 503	618 641 599	547 . 511 486	631 561 523	13.7 13.0 12.0	12.9 12.4 11.6	13.4 13.0 11.7	12.4 12.8 12.0	13.9 12.9 12.3	13.4 11.9				

Table 3. CHARACTERISTICS OF FARM AND NONFARM FAMILIES, BY RACE: 1976

(For meaning of symbols, see text)

	-	All races	}	•	White	•	, Black a	and other	races
Characteristic	Total	* Farm	Nonfarm	Total	Farm	Nonfarm	Total.	· Farm	Nonfarm
Total rimiliesthousands	56,245	2,200	54,045	49,873	2,105	47,768	% ,372	95	6,277
Metropolitan thousands.	37,801	416	37,386	32.848	403	32,446	4,953	13	4,940
Percent	67.2	18.9	69.2	65.9	19.1'	67.9	77.7	13.7	₹8.
Nonmetropolitan.	18,443	1,784	16,659	17,025	1,702	15,323	1,418	· 82	1,33
Percent	32,8	81.1	. 30.8	34.1	80.9	32.1	22.3	86.3	21.
Louis Cotton 1	7-1-		";"	-				j	
ill typės	100.0	100.0	100.0	100.0	100.0	100,0	100.0	100.0	100.
Husband-wife	84.1	92.2	83.8	86.8	92.8	86.6	62.9	78.9	62.
Male head, no wife present	2.6	3.5	2.5	2.4	3.2	2.3	4.1	10.5	÷ ·4.
Female head, no husband present	1-3.3	4.3	13.7	10.8	4.0	11.1	33.0	10.5	33.
remare mondy no made and processor try to		• • • • • • • • • • • • • • • • • • • •	21-		•	ţ	,)	
11 sizes	100.0	100.0	100.0	1 00.0	100.0	100.0	100.0	100,0	100
2 persons	37.8	37.9	37.8	38.8	38∌7	38.8	30.6	22.1.	30
3 to 5 persons.	52-8	50.4	52.9	52.9	50.5 يوس	53.0	51,8	48.4	. 51
6 or more persons	9.4	11.6	9.3	8.3	50.5 10.8	8.2	17.6	28.4	17
Le la	,				, ,,				
ean size of family	3.39	3.49	3.39	3.32	3.44	-3.32	3,92	4.56	. · ŝ.
ean size of lamily,				` `	٠	1.			-
11 families with own children under 18.	30,177	1,032	29,146	2 214	₹ 985	25,229.	3,963	4.7	3,9
Percent with-		i '	• • •	, -	} `}				•
1 or 2 own children under 18	70.8	66.9	A.0	72.0	4 68.0	. 72.2	63.2	(B)	63
3 or 4 own children under 18	24.6	123,8	24.6	24.2	22.9	24.2	27.1	,(B)°	26
5 or more own children under 18	4.0			3.8	9.0	3.6	9.7	(B)	9
y or more own carrains and a second	· · · · · •		Γ	ļ*	1 6	1	ł	, .	
ean number of own children.	2.04	2.19	2.04	2.00	2.17	1.99	2:34	(B)	2.
ean named of own carratements			1 .			1 '			
ercent of all families with members	•		1 '	•	-	ļ	`	,	
Under 18 years	55.8	49.0	56.1	54.1	48.3	54.3	69.1	65.3	69
18 to 64 years	92.4	89.9	92.5	92.0	89.8	92.1	95.5	91.6	95
65 years and over	17.7	24.4	17.4	18.0	23.8	17.7	15.2	36.8	14
oj jeara and over	15		1		<u> </u>		<u> </u>		

Source: Data from March 1976 Current Population Survey. See Current Population Reports, Series Pa20, No. 311, "Household and Family Sharacteristics: March' 1976." See tables A-1, A-2, and A-9 for standard errors.

Table 4. EMPLOYMENT STATUS OF THE FARM POPULATION 14 YEARS OLD AND OVER, BY SEX, APRIL 1976 AND 1970, AND BY REGION. APRIL 1976

(Numbers in thousands. Figures are five-quarter centered on April)

			North	-	, F	ercent d	istribution	
Sex and employment status	United	States	and, West	'South,	United	States	North and West	South
	1976	1970	1976	1976	1976	1970	1976	1976
3					- 5		100.0	1600.0
Both sexes	6,577	7,222	4,084	2,493	100.0	100.0	61.9	58.3
In labor force	3,980	4,293	2,526	1,454	60.5	59.4 40.6	38.1	41.7
Not. in labor Torce	2,597	2,929	1,557	1,040	34.5	40.0,	, ,,,,,	72.
, , , , , , , , , , , , , , , , , , , ,	3,980	4,293	2,526	1,454	100.0	100.Ó	100.0	100.0
In labor force	3,871	4,211	2,468	1,403	97.3	. 98.4	, 97.7	. 96.5
Employed	1,960	. 2,333	1,356	604	49.2	54.3	53.7	41.5
Agriculture		1.878	1,113	799	48.0	43.7	44.1	. 55.0
Nonagricultural industries	1,912	82	-, \$8	,51	2.7	1.9	2.3	3.5
Unemployed				,		*	100.0	100.0
/w->-	3,427	3,730	2,150	1,277	* 100.0	100.0	81.7	76.0
Male	2,727	2,974	1,756	, 971	79.6	79.7	18.3	24.0
Not in labor force	699	, 756 _°	393	306	20.4	20.3	,10.3	24.0
NOT IN TABOL CLOSES				971	100.9	100.0	100.0	100.0
In labor force	2,727	2,974	1,756	948	98.2	- 98.6	98.5	97.6
Employed	2,678	2,932	1,730	502	50.3	64.0	63.6	51.
Agriculture.	1,010	1,902	1,116	446	38.9	34.6	35.0	45.9
Nonagricultural industries	1,000	1,030	26	23	1.8	1.4	1.5	-, 2.
Unemployed	49	42] 20			1	↓ .	ļ .
•	1	3,492	1,934	1,216	100.0	100.0	100.0	100.
Female		1.319	770	, ,	30.7	37.8		39.
In labor force	1 '	• 2,173	1		60.3	62.2	60.2	60.
Not in labor force	1,090	1.	1,20				1	
	_	1,319	770	482	100.0	100.0	I	100.
In labor force.	1,193	1,279	738	455			4	94.
Employed		431	- 240	102			Y	21.
Agriculture	851	1	498			1 4		73.
Non-agricultural industries		40	32	27	4.7	1 3.0	4.2	5.

Table 5. EMPLOYMENT STATUS OF THE FARM POPULATION 14 YEARS OLD AND OVER, BY RACE, SEX, AND FOR REGION: APRIL (1976

(Numbers in thousands. Figures are five-quarter averages centered on April. For meaning of symbols, see text)

			,	Perc	ent distribu	tion
Labor force status, race, and sex	United States	North and	South	United States	North and	
	-		Journ.	States	West	South
WHITE		ĺ .		•		
Both sexes	6,170	4,045	2,125	100.0	100.0	100.0
In labor force	3,761	506ءُ 2	1,255	61.0	62.0	59.1
Not in labor force	2,409	1,540	869	39.0	38.1	40.9
In labor force	3,761	2,506	1,255	100.0	100.0	100.0
Employed	3,669	. 2,449	1,220	97.6	97.7	97.2
Agriculture	1,865/	1,341	524	49.6	* \$6.5	41.8
Nonagricultural industries	1,803	1,107	696	47.9	44.2	55.5
Unemployed	92	.57	_ 35	2.4.	2.3	2.8
. Male	3,214	2,128	1.086	100.0	100.0	100.0
In labor force	2,585	1,739	846	100.0 80.4	100.0	100.0
Not in labor force	628	389	239	19.5	18.3	·77.9 22.0
In labor force	2 505		`•			
Employed	2,585 2,544	1,739	846	100.0	100.0	100.0
Agriculture	1,539	1,713 1,103	831 d 436	98.4	98.5	98.2
Nonagricultural industries	1,006	610	396	59.5 38.9	63.4	51.5
Unemployed	41	26	15*	1.6	35.1 1.5	46.8 1.8
Female	2.056			5		
In labor force	2,956	1,917	1,039	100.0	100.0	100.0
Not in labor force	1,175	766 .1,152	629	39.7	40.0	39.4
	,,,,,,,,	.1,152	023	60.3	60.1	60,5
In labor force	1,175	766	409	100.0	100.0	100.0
Employed	124 ۽ 1	735	- 389	95.7	96.0	95.1
Agriculture	327	239	88	4 27.8	31.2	21.5
Unemployed	798 51	→ ⁴⁹⁷ / ₃₁	301	67.9	64.9	73.6
BLACK AND OTHER RACES	•	.,	, ,	4:3	4.0	4.9
BLACK AND OTHER RACES		1	أحساما	-	1	
Both sexes	. 407	378	369	100.0	(B)	100.0
In labor force.	21.9	21	_ 198	53.8	(B)	53.7
Not in labor force.	139	. 18	171	46.4	(B)	46.3
In labor force.			Į	Ī		
Employed	219	21	198	100.0	(B)	100.0
Agr#culture	203	20 . 14	183	92.7	" (B)	بقر 92
Nonagricultural industries.	108	, 385	103	42.9	(B)	49.4
(Unemployed)	. 16	- 1	15	. 49.3 7.3	(B) (B)	52.0 7.6
Mala				.		
In labor force	213	, 21	192	100.0	(B)	100.0
Not in labor force.	. 71	17	125	66.7	(B)	65.1
	Æ	,	, 66 L	33.3	(B)	34.4
In labor force:	1,42	√17	125	100.0	(B)	100.0
Employed.	134	· , 127	€ ',117	94.4	(B) 🛊	93.6
Monagricultural industries.	79	13	66	55.6	(B)	52,8
Unemployed	55 • 8	. 4	51	38.7	(B)	40.8
	<i>y</i> 1	, , -	` "	5.6	(B)	6.4
remale	194	17	177	100.0	(B)	.100.0
In labor force	77 '	5	72	39.7	(B)	40.7
4)	ا ز،	, 13	105	60.8	(B)	59.3
n fabor force	77 🏲	. 5	72	100.0	(B)	(B)
Riployed	. 69	4	65 `	89.6	(B)	(B)
Agriculture	15	12	13	19.5	(B) .	(B)
Unemployed	. 54	$\begin{bmatrix} 2 \\ 1 \end{bmatrix}$	· ′52	79-1	(B)	(B)
_ M	0 1		' 17	16.4	(15)	(B)

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Table 6. FARM RESIDENTS 14 YEARS OLD AND OVER EMPLOYED A AGRICULTURE BY CLASS OF WORKER, RACE, AND SEX, APRIL 1976 AND 1970, AND BY REGION: APRIL 1976

(Numbers in thousands: Figures are five-quarter averages centered on April. For meaning of symbols, see text)

		•			·\	Percent di	tribution	-	
	<u>.</u> *		North and	٠. ٠		•	North and	,	
Race, sex, and class of worker	Uni	ted St	ş West	South	Unt	ted States	West		South
· · · · · · · · · · · · · · · · · ·	1976	1970	1976	1976	1976	1970	1976	. 3	1976
TOTAL AGRICULTURAL WORKERS	·	:	• .		, ,	•		,	7
Both/sexes	. , 1,960	· · · 2; 333	1,356	604	100.0	100.0	100.0		100.4
San f-employed workers	1,210	1,411	841	₽ 369	61.7	60,5	62.0	Ì	61.1
When and salary workers	379	.395		145.	19.3	16.9	17.3	İ	24.
Unpaid family workers	370	. 4 526	281	89	18.9	, 22.5	20.7		14.
	1,618	1,902	1,116	502	100.0	. 1ρ0:0	100.0	;	100:0
Male	1,134	11,352	1	341	70,1	. 71.1	71.1	1	67.9
Sell-employed Workers.	330	349	204	* 126	20.4	18.3	18.3	•	25.
Wage and salary workers	153	, 200	118	35	9.5	10,5	. 10.6	١.	اء, 7
Oubsid family works	2		. *			•			
Female	342.	431	240	, 103	100.0	100.0	100.0	1	100.
	77	. 59	, 149	' ₩ 28	22.5	13:7) 20.4	į	27.
Self-employed workers	. 77	46		19,	14.3	10.7	12.5		18.
Unneid family workers	216	326	162	54	63/2.	75:6	, ,67.5	\	52.
Unpaid Pamily workers	~		· · · · ·	>	1	١.	' -		
WHITE								1	•
		ه ا	1 12/1	524	100.0	100.0	100.0	: _	100.
Both sexes	1,865	2,158	1,341	346	63:	62.9	,62.3	l	.66
Self-employed workers	1,182	1,358	225	105	17.7	13.9	16.8	i	20.
Wage and Balary workers	330			73	18.9	23,2	•20.9		13.
Unpaid family workers	2 53	501	, , , , , , , ,	1 , 72	10.7	1.		1	
	1,539	1,762	1,103	- 436	. 100.0	1 100.0	100.	1	100.
Male. J	1,106	1,304	787	319	• 71.9	74.0	71.4	•	73.
Self-employed workers	288	271	19			15.4	17.9	,	20.
Wage and salary workers	144	271 187	118			10.6	10.7	:	6.
unpaid family, workers	[/·:. ·····]	1				d			
Female	327	. 396	-239	88 1		100.0		1	100
Self-employed workers		54	49	11 . 27		13.6	1 , _	1	/ 30.
Wage and salary workers		28	28		7 •	7.1		1	15
Unpaid family workers	. 209	′ , 314	162	47	63.9	79.3	67.8	-	53.
			1 1	A	ļ	ļ. · . · .		}	•
BLACK AND OTHER RACES		1 .		100	· , · .	•,	,	-	•
Both sexes	94	17.	14	80	100.0	100.0	-(B)		100
	, .	7.				30.3	(B)	1	28
Self-employed workers	49	b & .		41	52.1	55.4	(B)	1 '	51
Unpaid family workers	138		4 8. *	16	19.1	·√r ₹ 14.3	(B)	1	20
Unpaid lamily workers	· •	٠.	* * * *	1 5	· * * * * * * * * * * * * * * * * * * *		(-)		,
Male	. 79					1		1:	(1
Self-employed workers	27	1 7 7				1 -		1.	. (
Wage and salary workers	. 42			35	53.2	9.3	1		. (
Unpaid family workers	10	1	3 1	3	12.7	ሳ · "·'	, ,,,,	.	
Years of the second	15	3.	· · · · · · · · · · · · · · · · · · ·	. 13	(B)	(B)	(B)		(
Pemale,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,	1		<u>.</u>	7		.			' (
Self-employed workers		_		_	, (B)	1 '		.	(
Wage and salary workers	a é	1 1	_		7 (B)	1 '	1 1 5		(1
Unpaid lamily workers,	۰-۱	-	- 1			1 .	1		

Table 7. FARM RESIDENTS 14 YEARS OLD AND OVER EMPLOYED IN NONAGRICULTURAL INDUSTRIES, BY CLASS OF WORKER, RACE, AND SEX, FOR REGIONS: APRIL 1976

(Numbers in thousands. Figures are five-quarter averages centered on April. For meaning of symbols, see text)

TOTAL NONAGRICULTURAL WORKERS Both sexes. Self-employed workers. Wage and salary workers. Unpaid family workers.	1,912 183 1,707	North and West	South	United States	North and	South
Both sexes	183 1, 707	1,113	., .			
Self-employed workers	183 1, 707	1.113	[•		
Self-employed workers	183 1, 707	1.113		,		
Wage and salary workers	1, 707		799	100.0	, 100.0	100.0
Unpaid family workers	1	98	. 85	9.6	8.8	10.6
		1,004	~ 703	89.3	90.2	88.0
Male	22	, 10	12	1.2	.9	1.5
	1.060	614	446	100.0	100.0	100.0
Self-employed workers	129	. 69	60	12.2	11.2	13.5
Wage and salary workers	928	• 544	384	87.5	88.6	86.1
Unpaid family workers	3		\ 2	3		.,
			` -			1 . 3.
Female	· /851	498	353	100.0	- 100.0	100.0
Self-employed workers	, 53	28	25	. 6.2	. 5.6	7.1
Wage and salary workers	779	460	319	91.5	92.4	90.4
Unpaid family workers	19	, 10	9	2.2	2.0	. 2.5
WHITE 5	•		-	,		,
*******	'		•	• '	_	,
Both sexes	.1,803	1,107	696	100.0	100.0	
Self-employed workers	177	• 97	1. 80	9.8	8.8	100.0
Wage and salary workers	1,604	999	605	- 89.0	90.2	86.9
Unpaid family workers	21	10	11	1.2	90.2	ئد أ. أ
Wale	1,006	*610	• 396	100.0	100.0	100.0
Self-employed workers	124	69 1	55	• 12.3	11.3	13.9
Wage and salary workers	878	540_	. 338	87.3	4 88.5	85.4
Unpaid family workers	3	1	, 2	. 3	ر. 2	, et .
Female	798	6 497		100.0	100 0	, ,,,,
Self-employed workers	53	28	301	100:0	100.0	100.0
Wage and salary workers	726	459	25 267	6.6 91.0	5.6	, 8.3
Unpaid family workers	18	9	* 20/	2.3	92.4	88.7
7	10		ا و	2.3	1.8	3.0
BLACK AND OTHER RACES	1	,		:	' †	
Both mexes	[,,,,		· 多世 秋		•	
	108	, 5	103	100.0.	(B)	100.0
Self-employed workers	. 5	- 1	. [4.6	(B)	4.9
Unpaid family workers	103	5	98	95:4	(B)	. , 95.1
***************************************		, , , , ,	, -		(B)	-
Male.,	• 55	4	- •51	. (B)	· (B)	(B)
Self-employed workers		, -1	5	, (B)	(B)	(B)
Wage and salary workers	50	4	46	(B)	(B)	(B)
Unpaid family workers	<u> </u>		-	(B)	(B)	, (B)
		!	, ,]	()	\ '	(2)
Female	54.	. ,2	52	(B)	(B)	, (B)
Self-employed workers		- 1		* (B))(B)	(B)
Wage and salary workers	53	. 1	. 52	" (B)	(B)	(B)
Unpaid family workers	- 1	. [-	K -	(B)	(B)	(B)

Appendix

DEFINITIONS AND EXPLANATIONS

Population coverage. With the exception of the total population shown in table. A, all figures in this report relate to the civilian noninstructional population. The total population shown in table B (210,332,000) differs from the estimated-April 1, 1976 total civilian population (2/12,611,000) chiefly in excluding the institutional population. For the Current Population Survey, both the institutional and military components of the population are regarded as entirely nonfarm.

Farm population. In the Current Population Survey, as in the 1960 and 1970 Censuses of Population, the farm population consists of all persons living in rural territory on places of 10 or more acres if as much as \$50 worth of agricultural products were sold from the place in the reporting year (for the CPS the preceding 12 months), it also includes those living on places of under 10 acres if as much as \$250 worth of agricultural products were sold from the place in the reporting year. Persons in institutions, summer camps, motels, and tourist camps, and those living on rented places where no land is used for farming, are classified as nonfarm

From April 1960 through January 1976, farm residence was determined in the Current Population Survey by the responses to two questions. Owners are asked, "Does this place have 10 or more acres?" and renters are asked, "Does the place you rent have 10 or more acres?" If the response is "Yes," the respondent is asked, "During the past 12 months, did sales of crops, livestock, and other farm products from this place amount to \$50 or more?" If the acreage response is "No," the inquiry relates to sales of \$250 or more

Beginning in February 1976, the second question was altered so that after responding either "Yes" or "No" to the acreage anguiry, owners/renters are asked, "During the past 12 months, how much did sales of crops, livestock and other farm products from this place amount to?" The respondents are given a choice of four answers "\$1,000 or more," "\$250 to \$999," "\$50 to \$249," and "Under \$50"

The question was changed to enable identification of the farm population as defined previously (see above) and as defined under the new farm definition announced by the U.S. Department of Agriculture and the Bureau of the Census in August 1975. Under the new definition, a farm is identified on the basis of sales alone, and is defined as any place from which \$1,000 or more of agricultural products are sold, or would probably be sold, from the place in the reporting year.

All of the farm figures presented in this report are based on the acreage-sales farm definition in use since-1960. Since

the \$1,000 sales item was not asked for all 5 of the months needed in computing the April-centered annual average for 1976 (see, "Five-quarter averages centered on April" in this section) comparable farm population estimates for both the old and new definitions are not available for 1976.

Farms located within the boundaries of urban territory, comprising a small minority of all farms, are not treated as farms for population census purposes, and their population is not included in the farm population. Urban territory includes all places with a population of 2,500 or more and the densely settled urbanized fringe areas around cities of 50,000 or more. Beginning with the 1972 estimate, the estimated farm population is limited to the rural territory as determined in the 1970 Census of Population In the Current Population Surveys of 1963 through 1971, the urban-rural boundaries used were those of the 1960 Census of Population and did not take into account the annexations and other substantial expansions of urban territory that were incorporated into the 1970 Census of Population The net effect was to classify an unknown number of persons as rural farm in the Current Population Surveys of 1970 and 1971 who were treated as urban (and hence nonfarm) in the 1970 census as well as in the Current Population Surveys beginning in 1972

In the Current Population Survey, inmarried persons attending college away from home are enumerated as resignents of their parents' homes, whereas in the Census of Population such persons are enumerated as residents of the communities in which they live while attending college. The effect of this difference is to classify a larger number of college-aged persons as farm residents in the Current Population Survey than would be so classified under decennial census usage.

Nonfarm population. The nonfarm population comprises, all persons living in urban areas and all tural persons not on farms / -

Five-quarter averages centered on April. April-centered annual averages of the farm population for the years 1970 through 1976 were computed by using data for the five-quarters centered on the April date for which the estimate was being prepared. For example, for April 1976; quarterly estimates for the months of October 1975, and January, April, July, and October 1976, were used with a weight of one-eighth given to each of the two October estimates and a weight of one-fourth to each of the estimates for the other 3 months. One reason, for the choice of April as the date for centering population estimates is that this is the decennial census month.



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April centered annual averages for persons under 14 years by race and sex, and for persons 14 years old and over, by race, sex, age, labor force characteristics, and region were also computed for 1976 by using data for the specified characteristics for the five quarters centered on April 1976.

Metropolitan-nonmetropolitan residence. The population residing in standard metropolitan statistical areas (SMSA's) constitutes the metropolitan population. The metropolitan population in this report is based on SMSA's as defined in the 1970 population census publications and does not indude any subsequent additions or changes. For the 1970 census, except in New England, an SMSA is a county or group of contiguous counties which contains at least one city of 50,000 inhabitants or mage, or "twin crities" with a combined population of at least 50,000. In addition to the county, or counties, containing such a city or cities, contiguous counties are included in an SMSA if, according to certain criteria, they are essentially metropolitan in character and are socially and economically integrated with the central county. In New England, SMSA's consist of towns and cities, rather than counties.

Geographic regions. The major regions of the United States for which data are presented represent groups of States, as follows:

North and West: Nor least, North Central, and West regions combined.

Northeast: Connecticut, Maine, Massachùsetts, New Hampshire, New Jersey, New York, Pennsylvania, Rhode Island, Vermont.

North Central: Illionois, Indiana, Iowa, Kansas, Michigan, Minnesota, Missouri, Nebraska, North Dakota, Ohio, South Dakota, Wisconsin.

West: Alaska, Arizona, California, Colorado, Hawaii, Idaho, Montana, Nevada, New Mexico, Oregon, Utah, Washington, Wyoming.

South: Alabama, Arkansas, Delaware, District of Columbia, Florida, Georgia, Kentucky, Louisiana, Maryland, Mississippi, North Carolina, Oklahoma, South Carolina, Tennessee, Texas, Virginia, West Virginia

Age. The age classification is based on the age of the person at last birthday.

Race. The population is divided into three groups on the basis of race: White, Black, and "other races." The last category includes Indians, Japanese, Chinese, and any other race except White and Black. In the text of this report, Blacks refer to Blacks and persons of races other than White.

Family. The term "family," as used in this report, refers to a group of two-or more persons related by blood, marriage, or adoption and residing together, all such persons are

considered as members of the same family. Thus, if the son of the head of the household and the son's wife are in the household, they are treated as part of the head's family. On the other hand, a lodger and his wife not related to the head of the household or an unrelated servant and his wife are considered as additional families, and not a part of the household head's family.

The mean size of family is derived by dividing the number of persons in families by the total number of families. In the classification of families by number of family members, the head of the family and all other persons in the family are included. The number of family members is the same as size of family.

Head of family. One person in each family was designated as the head. The head of a family is usually the person regarded as the head by members of the family. Women are not classified as heads if their husbands are resident members of the family at the time of the survey. Married couples related to the head of a family are included in the head's family and are not classified as separate families.

The Census Bureau has traditionally designated a head of household to serve as the central reference person for the collection and tabulation of data for individual members of the household (or family). However, recent social changes have resulted in a trend toward more equal status for all members of the household (or family), making the term "head" less relevant in the analysis of household and family data. As a result, the Bureau is currently developing new techniques of enumeration and data presentation which will eliminate the concept of "head." While some of the data in this report are based on the concept of "head," methodology for future Census Bureau reports will reflect a gradual movement away from this traditional practice

Type of family. The classification of families by type is based on the sex and marital status of head. Families with a head and wife present are termed "husband wife" families. Families in which the spouse of the head is not present are families with "other male head" or "female head" depending on the sex of the head.

Own children. "Own" children in a family are single (never precised) sons and daughters, including stepchildren and adology children, of the family head. In table 3, the mean number of own children is derived by dividing the number of children by the total number of families with own children under 18.

Marital status. Data refer to present marital status. The primary categories of marital status are single (never married) and ever married. The following sub-categories of ever married may be distinguished: (1) married, spouse present; (2) married, spouse absent (excluding separated); (3) separated; (4) yudowed; or (5) divorced.

Lifetime birth expectations. Lifetime births expected are determined by adding any additional births a woman expects

to the children she has already borne, if any. Questions regarding expected additional births were asked in June 1976 of women 14 to 39 years old who were currently married (spouse present or spouse absent excluding separated).

Births to date. In the data on birth expectations of wives in table C, the number of "births to date," has the same meaning as the number of children ever born.

Children ever born. The term "children ever born" refers to the total number of live births reported by ever-married women. Included in the number are children born to the woman before her present marriage, children no longer living, and children away from home, as well as children who were still living in the home.

Labor force and employment status. The definitions of labor force and employment status in this report relate to the population 14 years old and over.

Labor force. Persons are classified as in the labor force of they were employed as civilians, unemployed, or in the Armed Forces during the survey week. The "civilian labor force" is comprised of all civilians classified as employed or unemployed.

Employed. Employed persons comprise (1) all civilians who, during the specified week, did any work at all as paid employees or in their own business or profession, or on their own farm, or who worked 15 hours or refere as unpaid workers on a farm or in a business operated by a member of the family, and (2) all those who were not working but who had jobs or businesses from which they were temporarily absent because of illness, bad weather, vacation, or labormanagement dispute, or because they were taking time off for personal reasons, whether or not they were paid by their employers for time off, and whether or not they were seeking other jobs. Excluded from the employed group are persons whose only activity consisted of work around the house (such as own home housework, painting or repairing own home, etc.) or volunteer work for religious, charitable, and similar-organizations.

Unemployed. Unemployed persons are those civilians who, during the survey week, had no employment but were available for work and (1) had engaged in any specific job-seeking activity within the past 4 weeks, such as registering at a public or private employment office, meeting with prospective employers, checking with friends or relatives, placing or answering advertisements, writing letters of application, or being on a union or professional register, (2) were waiting to be called back to a job from which they had been laid off, or (3) were waiting to report to a new wage or salary job within 30 days.

Not in the labor force. All civilians who are not classified as employed or unemployed are defined as "not in the

labor force." This group who are neither employed nor seeking work includes persons engaged only in own home housework, attending school, or unable to work because of long-term physical or mental illness; persons who are retired or too old to work; seasonal workers for whom the survey week fell m an off season; and the voluntarily idle. Persons doing only unpaid family work (less than 15 hours) are also classified as not in the labor force.

Agriculture. The industry category "agriculture" is somewhat more inclusive than the total of the two major occupation groups, "farmers and farm managers" and "farm laborers and supervisors." It also includes (1) persons employed on farms in occupations such as truck driver, mechanic, and bookkeeper, and (2) persons engaged in certain activities other than strictly farm operation such as cotton ginning, contract farm services, veterinary and breeding services, hatcheries, experimental stations, greenhouses, landscape gardening tree service, trapping, hunting preserves, and kennels.

Nonagricultural industries. This category includes all industries not specifically classed under agriculture.

Multiple jobs. Persons with two or more jobs during the survey week were classified as employed in the industry in which they worked the greatest number of hours during the week. Consequently, some of the persons shown in this report as engaged in nonagricultural activities also engaged in agriculture and vice versa.

Class of Worker

Self-employed workers. Persons who worked for profit or fees in their own business, profession, or trade, or who operated a farm either as an owner or tenant.

Wage and salary workers. Persons who worked for any governmental unit or private employer for wages, salary, commission, tips, pay "in kind," or at piece rates.

Unpaid family workers. Persons who worked without pay on a farm or in a business operated by a person to whom they are related by blood or marriage.

Income. Total money income is the algebraic sum of the amounts received in the preceding calendar year from each of the following sources: (1) Money wages or salary; (2) net income from nonfarm self-employment, (3) net income from farm self-employment; (4) Social Security or railroad retirement; (5) dividends, interest (on savings or bonds), income from estates or trusts, or net rental income; (6) public assistance or welfare payments; (7) unemployment and workmen's compensation, government employee pensions, or veterans' payments; (8) private pensions, annuities, alimony, regular contributions from persons not living in this household, and other periodic, income.



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Receipts from the following sources are not included as income: (1) Money received from the sale of property, such as stocks, bonds, a house, or a car (unless the person was engaged in the business of selling such property, in which case the net proceeds would be counted as income from self-employment); (2) withdrawals of bank deposits, (3) money borrowed; (4) tax refunds; (5) gifts; and (6) lumpsum inheritances or insurance payments.

Family income. The total income of a family is the algebraic sum of the amounts received by all income recipients in the family.

In the income distribution for families, the lowest income group (less than \$4,000), includes those families who were classified as having no income in the income year and those reporting a loss in net income from farm and nonfarm self-employment or in rental income. Many of these were living on income "in kind," savings, or gifts, or were newly constituted families, or families in which the sole breadwinner had recently died or had left the household. However, many of the families who reported no income probably had some money income which was not recorded in the survey.

It should be noted that although the income statistics refer to receipts during the preceding year, the composition of families refers to the time of the survey. The income of the family does not include amounts received by persons who were members of the family during all or part of the income year if these persons no longer resided with the family at the time of enumeration. On the other hand, family income includes amounts reported by related persons who did not reside with the family during the income year but who were members of the family at the time of enumeration.

The median income is the amount which divides the distribution into two equal groups, one having incomes above the median, and the other having incomes below the median. The medians for families are based on all families.

Low-income (poverty) definition. Families and unrelated individuals are classified as being above or below the low-income level using the poverty index adopted by a Federal Interagency Committee in 1969. This index is based on the Department of Agriculture's 1961 Economy Food Plan and reflects the different consumption requirements of families based on their size and composition, sex and age of the family head, and farm nonfarm residence. In order to keep the poverty index constant over time, the thresholds are updated annually based on changes in the Consumer Price Index. The low-income threshold for a nonfarm family of four was \$5,500 in 1975, \$4,275 in 1972, and \$2,973 in 1959. Corresponding low-income thresholds for a farm family of four were \$4,695 in 1975, \$3,643 in 1972, and \$2,539 in 1959.

In analyzing data on the low-income population, the following limitations should be noted. The low-income concept has been developed in order to identify, in dollar terms, a minimum level of income adequacy for families of different

types in keeping with American consumption patterns. Based on an analysis of the percent of income devoted to food expenditures, an estimate was developed of the minimum cost at which an American family, making average choices, can be provided with a diet meeting recommended nutritional goals. Consequently, it is an overall statistical yardstick which reflects the different consumption requirements of families of different size, taking into account family composition and farm nonfarm residence. Insofar as individual circumstances or consumption patterns differ, the dollar value of the low income threshold for a given family size may not represent the money income required by an individual family to maintain a level of economic well-being equivalent to other families with similar incomes.

Average annual rate of change. Average annual rates of change are estimated using an exponential model. Specifically the average annual rate of change is defined to be the value x which satisfies the relationship

$$y_t = y_0 e^{tx}$$
 (1)

Here y_0 is the population at an initial point in time, \dot{y}_t is the population at a later point in time, and t is the number of years that have elapsed between the measurements of the population sizes. The annual average rate of change is estimated using the estimates of the population for the two points in time in equation (1) and solving for x.

Symbols. A dash "-" represents zero and the symbol "B" means that the base for the derived figure is less than 75,000.

Rounding. The individual frgures in this report are rounded to the nearest thousand. With few exceptions, the individual figures have not been adjusted to group totals, which are independently rounded. Percentages are rounded to the nearest tenth of a percent; therefore, the percentages in a distribution do not always add to exactly 100.0 percent. The totals, however, are always shown as 100.0. Percentages are, based on the rounded absolute numbers.

SOURCE AND RELIABILITY OF THE ESTIMATES

Source of data. Most of the estimates in this report are April centered five-quarter averages of data collected in 1960 through 1976 from the Current Population Survey (CPS) of the Bureau of the Census. The monthly CPS deals mainly with labor force data for the civilian, noninstitutional population. Questions relating to labor force participation are asked about each member 14 years old and older in each sample household. Data on fertility and birth expectations, income and low income status for the year 1975, and household and family characteristics of farm and nonfarm families are not based on five-quarter averages. These types of data are obtained from supplementary questions to CPS



asked in the months of June (fertility) and March 1976. The farm and nonfarm residence data for persons are April-centered five-quarter averages.

The present CPS sample was ffitially selected from the 1970 census file and is updated continuously to reflect new construction where possible (see section "Nonsampling Variability" below). Previous sample designs used, as a basis, files from the census most recently completed at the time.

The following table provides a description of some aspects of the CPS sample designs in use during the referenced data collection periods

The estimation procedure used for the monthly CPS data involves the inflation of the weighted same results to independent estimates of the civilian noninstitutional population of the United States by age, race, and sex. These in-

Description of the Current Population Survey

		· Household	eligible ,	Households	
Time period	Number of sample areas	Intervièwed	Not intexviewed	visited, not eligible ²	
March 1977 ³	*614 *461 449	53,500 45,000 48,000	2 ,5 00 2,000 2,000	9,500 8,000 8,500	

These areas were chosen to provide coverage in each State and the District of Columbia.

These are households which were visited but were found to be vacant or otherwise not eligible for interview.

³A supplementary sample of housing units in 24 States and the District of Columbia was incorporated with the monthly CPS to produce March 1977 data.

dependent estimates were based on statistics from decennial censuses; statistics on births, deaths, immigration, and emigration; and statistics on the strength of the Armed Forces. For estimates from March CPS data in this report persons in the Armed Forces were also included, and the estimation procedure in March for data in this report also involves a further adjustment so that husband and wife of a household receive the same weight.

Decennial census of population. Decennial census data in this report are based on complete counts of on the samples associated with the census as indicated in the list of sources. Descriptions of samples from the census are found in the appropriate census publications. To determine if the 1950, 1960, and 1970 data in the text tables of this report are based on complete counts or on the samples associated with the census refer to the sources of data at the bottom of that table.

Reliability of the estimates. Since the estimates in these tables were based on a sample, they may differ somewhat from the figures that would have been obtained if a complete census had been taken using the same schedules, instructions, and enumerators. There are two types of errors possible in an estimate based on a sample survey—sampling and nonsampling. The standard errors provided for this report primarily indicate the magnitude of the sampling error. They also partially measure the effect of some nonsampling errors in response and enumeration, but do not measure any systematic biases in the data. The full extent of nonsampling error is

unknown. Consequently, particular care should be exercised in the interpretation of figures based on a relatively small number of cases or on small differences between estimates.

Nonsampling variability. As in any survey work, the results are subject to errors of response and nonreporting in addition to sampling variability. Nonsampling errors can be attributed to many sources, e.g., inability to obtain information about all tases in the sample, definitional difficulties, differences in the interpretation of questions, inability or unwillingness to provide correct information on the part of respondents, inability to recall information, mistakes made in collection, such as in recording or coding the data, mistakes made in processing the data, mistakes made in estimating values for missing data, and failure to represent all units with the sample (undercoverage).

The approximate magnitude of three sources of undercoverage in CPS is known. About 600,000 conventional new
construction units (housing units, other than mobile homes
or group quarters) were issued building permits prior to the
1970 census but building was not completed by the time of
the census (i.e., April 1970), these units have no representation in the CPS sample. Most conventional new construction
for which building permits were issued after the census, is
represented. About 290,000 occupied mobile homes are not
represented in CPS; these units were either missed in the
census or have been built or occupied since the census. In
addition, about 30,000 units are not represented either because they have been converted from nonresidential units or



are houses moved to different sites since the census. The extent of other sources of undercoverage is unknown.

Note: These estimates of missed units are relevant to the present sample only and not to earlier designs where the extent of undercoverage was generally less.

In most cases the schedule entities for income are based on the memory or knowledge of one person, usually the wife of the family head. The memory factor in data derived from field surveys of income probably produces underestimates because the tendency is to forget minor or irregular sources of income. Other errors of reporting are due to misrepresentation or to misunderstanding as to the scope of the income concept.

Sampling variability. The standard errors given in the following tables are primarily measures of sampling variability, that is, of the variations that occurred by chance because a sample rather than the whole of the population was surveyed. The chances are about 68 out of 100 that an estimate from the survey differs from a complete census figure by less than the standard error. The chances are about 90 out of 100 that this difference would be less than 1.6 times the standard error and about 95 out of 100 that the difference would be less than twice the standard error.

All of the statements of comparison appearing in the text are significant at a 1.6 standard error level or better, and most are significant at a level of more than 2.0 standard errors. This means that for most differences cited in the text, the estimated difference is greater than twice the standard error of the difference. Statements of comparison qualified in some way (e.g., by use of the phrase, "some eyidence") have a level of significance between 1.6 and 2.0 standard errors.

shown in the report only when the base is 75,000 or greater. Because of the large standard errors involved, there is little chance that percentages would reveal useful information when computed on a smaller base. Estimated numbers are shown, however, even though the relative standard errors of these numbers are larger than those for corresponding persentages. These smaller estimates are provided primarily to permit such combinations of the categories as serve each ser's needs.

Comparability with other data. Data obtained from the CPS and other sources are not entirely comparable. This is due in large part to differences in interviewer training and experience and in differing collection procedures. These differences are not reflected in the standard errors provided. Therefore, caution should be used in comparing results between different sources. See the appendix for more details on the comparability of CPS and other data.

Standard errors for data based on the decennial census. Sampling errors of all data from the samples of the decennial

censuses shown in this report except for fertility are small enough to be disregarded. The standard errors for census sample data may be found in the appropriate census volumes.

Standard error tables and their use. In order to derive standard error that would be applicable to a large number of estimates and could be prepared at a moderate cost, a number of approximations were required. Therefore instead of providing a standard error for each estimate, generalized sets of standard errors are provided for various types of characteristics. As a result the sets of standard errors provided give an indication of the order of magnitude of the standard error of an estimate rather than the precise standard error.

gures presented in tables A-1, A-2, and A-3 provide approximations to standard errors of various estimates for families, unrelated individuals, and persons. Table A-4 provides approximations to the standard errors of estimated fertility rates for the nonfarm population. Estimated standard errors cannot be obtained from tables A-1, A-2, and A-3 without the use of the factors in table A-5. These factors must be applied to the generalized standard errors in order to adjust for the combined effect of sample design and the estimating procedure on the value of the characteristic. The standard error tables with which each factor should be used are indicated in table A-5. Standard errors for intermediate values not shown in the generalized tables estandard errors may be approximated by interpolation

Table A-1. Standard Errors of Estimated Numbers of Persons or Families in the Farm Population

(08 Chances out of 100. Numbers continues and							
Size of estimate	Standard error						
25:	6 9 13 20 29 42 70 107 173 235						

Note: For standard errors for metropolitan or nonmetropolitan data multiply the standard errors above by 1.4.

For standard errors for the years 1960 to 1966 multiply the above standard errors by 1.2

^{**} Estimated standard errors cannot be calculated for table C (fertility rates) without the use of the bases in table A-6

Two parameters are used (denoted as "a" and "b") to calculate standard errors for each type of characteristic, they are presented in table A-5. These parameters were used to calculate the tabulated standard errors in tables A-1, A-2, and A-3 and to calculate the factors in table A-5. They also may be used to calculate the standard errors for estimated numbers and estimated percentages directly. Methods for direct computation are given in the following sections.

Standard errors of estimated numbers. The approximate standard error, $\sigma_{\rm X}$ of an estimated number shown in this report can be obtained in two ways. It may be obtained by use of the formula

$$\sigma_{\mathbf{v}} \stackrel{\circ}{=} \mathbf{f} \sigma$$
 (1)

where f is the appropriate factor from table A.5, and d is the standard error on the estimate obtained by interpolation from table A.1 or A.2. Alternately, standard errors may be approximated by the following formula, (2), from which the standard errors were calculated in tables A.1 and A.2. Use of this formula will provide more accurate results than the use of formula (1) above.

$$\sigma_{\mathbf{x}} = \sqrt{\mathbf{a}\mathbf{x}^2 + \mathbf{b}\mathbf{x}} \tag{2}$$

Here x is the size of the estimate and a and b are the parameters in table A-5 associated with the particular type of characteristic.

Table -2. Standard Errors of Estimated Numbers of Persons or Families in the Total or Non-farm Population

(68 chances out of 100. Numbers in thousands)

Standard error
6 9 13 20 28 40 63 88 122 145 180 223
199

Note: For standard errors for metropolitan or nonmetropolitan data multiply the standard errors above by 1.4.

For standard errors for the years 1960 to 1966 multiply the above standard errors by 1.2

Standard errors of estimated percentages. The reliability of an estimated percentage, computed using sample data for both numerator and denominator, depends on both the size of the percentage and the size of the total, upon which this percentage is based. Estimated percentages are relatively more reliable than the corresponding estimates of the numerators of the percentages, particularly if the percentages are 50 percent or more. When the numerator and denominator of the percentage are in different categories, use the factor or parameters indicated by the numerator. The approximate standard error, $\sigma_{(\mathbf{x},\mathbf{p})}$, of an estimated percentage can be obtained by use of the formula

$$\sigma_{(\mathbf{x},\mathbf{p})} = \mathbf{f}\sigma$$
 (3)

In this formula fis the appropriate factor from table A-5 and σ is the standard error on the estimate from table A-3. Alternately, standard errors may be approximated by the following formula, (4), from which the standard errors in table A-3 were calculated, direct computation will give more accurate results than use of the standard error tables and the factors.

$$\sigma_{(\mathbf{x},\mathbf{p})} = \sqrt{\frac{\mathbf{b}}{\mathbf{x}} \cdot \mathbf{p} (100 \cdot \mathbf{p})} \tag{4}$$

Here x is the size of the subclass of persons or families and unrelated individuals which is the base of the percentage, p is the percentage (0 \leq p \leq 100), and b is the parameter in table A-5 associated with the particular type of characteristic in the numerator of the percentage.

Illustration of use of standard error tables. Table D of this report shows that in 1976 there were 3,592,000 persons employed in agriculture. Table A-5 shows that the appropriate factor is 0.9 and that this factor is to be used with the standard errors in table A-1. Table A-1 shows the standard error on an estimate of this size to be approximately 86,200. Applying the factor of 0.9 and using formula (1), the approximate standard error is 0.9 x 86,200 = 78,000. The chances are 68 out of 100 that the estimate would have been a figure differing from a complete census figure by less than 78,000. The chances are 95 out of 100 that the estimate would have differed from a complete census figure by less than 156,000 (twice the standard error).

Of these 3,592,000 persons employed in agriculture 2,941,000 or 81.9 percent are males. From table A-5 the appropriate b parameter for computing standard errors is 1334.7957; using formula (4), the standard error on an estimate of 81.9 percent is

$$\sqrt{\frac{1334.7957}{3,592,000}}$$
,81.9 (190-81.9) ≈ 0.7 percent

² Formula (2) gives a standard error of 84,000.

Consequently, chances are 68 out of 100 that the estimated 81.9 percent would be within 0.7 percentage points of a complete census figure. Chances are 95 out of 100 that the estimates would be within 1.4 percentage points of a complete census figure, i.e., the 95-percent confidence interval would be from 80.5 to 83.3 percent.

*Standard error of a difference. For a difference between two sample estimates, the standard error is approximately equal to

$$\sigma_{(\mathbf{x}\cdot\mathbf{y})} = \sqrt{\sigma_{\mathbf{x}}^2 + \sigma_{\mathbf{y}}^2} \tag{5}$$

Table A-3. Standard Errors of Estimated Percentages

(68 chances out of 100)

Base of percentages	-		Estimated p	ercentage	•		
(thousands)	. 1 or 99	2 or 98	5 or 9.5 x	10 or 90	25 or 75 ·	, 50	
5,	. 2.5	3.5	5.5	7.6.	10.9	- 12.0	
0	1.8	2.5,	3.9	5.4	7.7	8.	
00	1.3	1.8	2.8	3.8	5.5	6.	
50	0.8	1.1 0.8	1.7	2.4	3.5	.4'. 2 .	
,000	0.4	0.6	0.9	1.2	1.7	2.	
,500	0.3	0.4 ′	0.6	0.8	1.1	1.	
,000	0.2	- 0.3	0.4	0.5	0.8	0.	
9,000	0.13	0.2	0.3	0.4	0.5	0. 0.	

Note: For metropolitan or nonmetropolitan standard errors, multiply appropriate standard errors above by 1.4. For standard errors for the years 1960 to 1966, multiply avove standard errors by 1.22.

Table A-4. Standard Errors of Estimated Fertility Rates for the Nonfarm Population

(68 chances out of 100)

Number of women	Children ever born per 1,000 women								
. (thousands)	′500	1,000	1,500	2,000	12,500	3,000	3;500.	4;000	
250	51 36 30 26 48 11 9 7 6 5	93 66 54 47 33 20 15 12 11 9 8	129 92 74 65 45 29 20 16 15 12	164 116 95 82 58 • 37 26 21 19 16 15	198 140 114 99 70 44 31 26 23 20 19	234 166 135 117 83 . 52 38 29 27 24 22 20	274 194 158 137 97 61 44 35 31 28 25 23	315 -222 181 158 112 -70 -50 41 35 32 29	

Note: Multiply these standard errors by 1.38 to obtain standard errors for fertility of the farm, population.



where σ_x and σ_y are the standard errors of the estimates x and y; the estimates can be of numbers, percents, averages, etc. This will represent the actual standard error quite accurately for the difference between two estimates of the same characteristic in two different areas, or for the difference between two separate and uncorrelated characteristics in the same area. If however, there is a high positive correlation between the two characteristics, the formula will overestimate the true standard error.

illustration of the computation of the standard error of a difference. Table 1 of this report shows that there were 4,305,000 males and 3,947,000 females on farms in 1976. The estimated difference between the number of males on farms and the number of females on farms is 358,000m.

Using formula (2) and the appropriate parameters from table A-5, the standard error on the estimate of 4,305,000 males on farms is 97,000.

Sintilarly the approximate standard error on the estimate of 3,947,000 females on farms is 92,000. Therefore, from formula (5) the approximate standard error on the estimate difference of 358,000 persons is

$$134,000 \pm \sqrt{(97,000)^2 + (92,000)^2}$$

This means the chances are 68 out of 100 that the estimated difference based on the sample estimates would vary from the difference derived using complete census figures by less than 134,000 persons. The 68 percent confidence interval about the 358,000 persons difference is from 224,000 to 492,000 i.e., 358,000 ± 134,000. A conclusion that the average estimate of the difference derived from all possible samples of the same size and design lies within a range computed in this way would be correct for roughly 68 percent of all possible samples. The 95-percent confidence interval is 90,000 to 626,000. Thus, we can conclude with 95-percent confidence that there was a significant difference in the numbers of males and females on farms in 1976.

Standard error of a ratio. Certain mean values for persons in families shown in the tables of this report were calculated as the ratio of two numbers. For example, the mean number of persons per family is calculated as

x total number of persons in families
y total number of families

Standard errors for these means may be approximated as shown below. There are two cases to consider. In either case, the denominator y represents a count of families of a certain.

class, and the numerator x represents a count of persons with the characteristic under consideration who are members of these families.

Case 1: There is at least one person having the characteristic in every family of the class: as for example, the mean number of persons per family or the mean number of persons per family with a male head. For the track of this kind, the standard error is approximated by the following formula:

$$\frac{\sigma_{x}}{y} = \sqrt{\left(\frac{x}{y}\right)^{2} \left[\left(\frac{\sigma_{y}}{y}\right)^{2} + \left(\frac{\sigma_{x}}{x}\right)^{2} - 2\rho\left(\frac{\sigma_{x}}{x}\right)\left(\frac{\sigma_{y}}{y}\right)\right]}, \quad (6)$$

The standard error of the estimated number of families, σ_y , and the standard error of the estimated number of persons with the characteristic in those families, σ_x , may be calculated by the methods described above. In formula (6), ρ represents the correlation coefficient between the numerator and the denominator of the estimate. In the above examples, and for other ratios of this kind, use 0.7 as an estimate of ρ .

Case 2: The number of persons having the characteristic in a given family may be 0, 1, 2, 3, or more: for example, the mean number of persons under 18 years of age. For ratios of this kind the standard error is approximated by formula (6) but ρ is assumed to be zero. If ρ is actually positive, then this procedure will provide an overestimate of the standard error of the ratio.

Standard error of a fertility ratio. Table A-4 provides standard errors for both number of children ever born and number of expected lifetime births per 1,000 women.³ The sampling variability on the ratio of children born per 1,000 women depends on the shape of the distribution on which the ratio is based, the size of the sample, the sample design and the use of ratio estimates.

Illustration of the computation of the standard error of a fertility ratio. Table C shows that in 1976 there were 2,123 children ever born per 1,000 ever married far favormen aged 25 to 34. Table A-6 shows that there were about 324,000 women in this group. Table A-4 shows the standard error of a

³The bases for the estimated fertility rates are given in table A-6 for use with table A-4 to obtain estimated standard errors.

Table A-5. Parameters or Factors and Standard Error Tables to be Used to Obtain Standard Errors for Each Type of Characteristic

	- Param	eters		ors and rror, tables
Type of characteristic	a	. в "	f factors	Standard error tables
PERSONS		*	,	
Farm population: Total, agriculture employed, or				• · · :
nonagriculture employed	0.00013886	1597.6160	. 1.0	A-1_A-3
Total or nonfarm population: Agriculture employment	0.000177	1334.77957	.0.9	A-1,A-3
Total or nonagriculture employment: Total or White	-0.000006	759.3218	0.7	A-2,A-3
Black and other races	-0.000049	680.0632	0.7	A-2,A-3
FAMILIES AND HOUSEHOLDS	-			,
Farm population: Total or White	0.000179	2652.3108	1.3	∕A-1,A-3
Black and other races	0.000662	.2397.1230	1.2	A-1,A-3
Total or nonfarm population: Total or White	-0.000010	1388.6444	≱ 0.9.	A-2,A-3
Black and other races	-0.000087	1255.0382	0.9	^ A-2,A-3
FAMILY INCOME	,			
Farm: Total or White	0.000143	2030.6775	1.1	- A-1,A+3
Total or White	0.001144	. 1761. 1516	1.0	A-1,A-3
Total or nonfarm: Total or White	-0.000008	1063.1809	. 0.8	A-2,A-3
Black and other races	-0.000064	922.0689	0.8	A-2,A-3
FERTILITY			,	
Number of women	0.000322	2993.0343	1.4	A-1,A-3
Total or nonfarm: Number of women	-0.000018	≱ 567.0337.	/ 1.0	A- 2,A-3

Note: For standard errors for metropolitan or nonmetropolitan data, multiply the appropriate parameter by 2.0. For standard errors for the years 1960 to 1966 multiply the appropriate parameter by 1.5.

24

1,767

rate of 2,123 children on a base of 324,000 women to be approximately 157. Multiplying the standard error of 157 by 1.38 (factor for fertility standard errors of the farm population), the standard error becomes 217. Consequently the chances are 68 out of 100 that the estimate would have, shown a fertility rate differing from a complete cense figure.

14 to 39 years old, τeporting

birth expectations.....

by less than 217. The chances are 95 out of 100 that the estimate would have shown a fertility rate differing from a complete census figure by less than 434 (twice the standard error), i.e., this 95-percent confidence interval would be between 1,689 and 2,557 children ever born per 1,000 ever-married farm women aged 25 to 34.

Table A-6. Estimates of the Number of Ever-Married Women and Number of Currently Married Women Reporting Birth Expectations, by Age, Race and Farm-Nonfarm Residence: June 1976 CPS

(Numbers in thousands) Black and other races White Total Women by age Total Farm Nonfarm Nonfarm Total Farm Nonfarm Total Farm. WOMEN EVER MARRIED 41' 3,856 3,897 31,002 28,010 864 27,146 31,907 905 Total, 15 to 44 years..... 5,766 9 **-719** 101 728 6,595 110 6,485 5,867 15 to 24 years..... 1,709 14 13,812 12,414 310 12,103 1,722 14,136 324 25 to 34 years.... 1,446 1,429 10,705 9,729 453 9,276 11,175 471 35 to) 44 years: .:.... WOMEN CURRENTLY MARRIED

Source: U.S. Bureau of the Census, Current Population Survey, June 1977.

19 267

Note: This table is to be used for computation of the standard errors of estimated fertility rates in table C.

18,742

524

500

17,476

16,975

1,791